

Mathematics 375 – Probability Theory
Solutions for Midterm Exam 2 – November 17, 2011

I. Birth weights of babies born in the US are normally distributed with mean $\mu = 3315$ grams and standard deviation $\sigma = 575$ grams.

- A) (5) Find the (approximate) probability that a baby will have birth weight more than 3700 grams.

Solution: Let $Y =$ birth weight of a randomly selected baby, so by the given information, $Y \sim Normal(3315, 575^2)$. Standardizing, we have

$$P(Y > 3700) = P\left(\frac{Y - 3315}{575} > \frac{3700 - 3315}{575}\right) \doteq P(Z > .67) \doteq .2514$$

(since $\frac{Y-3315}{575} = Z \sim Normal(0, 1)$).

- B) (10) If 20 newborn babies are selected randomly, how many of them would be expected to have weights between 2740 and 3890 grams?

Solution: We have $2740 = 3315 - 575 = \mu - \sigma$ and $3890 = \mu + \sigma$. The probability that a single baby has a weight in this range is

$$P(2740 < Y < 3890) = P(-1 < Z < 1) = 1 - 2P(Z > 1) = 1 - 2(.1587) = .6826$$

(Note that this interval is where the 68% in the Empirical Rule comes from.) If X is the number of babies in the 20 that have weights in this range, then technically X would have a *hypergeometric* distribution, since the selection *would not be done with replacement*. However, we don't know the size of the entire population of babies in question. In any case that would be much larger than 20. So the *binomial* formulas can be used to estimate this. Say $X \sim Binomial(20, .6826)$. Then

$$E(X) = n \cdot p = 20 \cdot 0.6826 = 13.652.$$

So we expect this many babies to have weights in this range.

- C) (10) If we knew the mean and standard deviation, but we did not know the precise distribution of the birth weights, what could we say in general about the probability that a randomly selected baby has weight differing from 3315 grams by at least 862.5 grams?

Solution: By Tchebysheff's Theorem, since $862.5 = 1.5 \cdot 575$

$$P(|Y - \mu| \geq 1.5\sigma) \leq \frac{1}{(1.5)^2} = .4444.$$

II. (15) The number of defects Y in a 1-foot segment of a magnetic tape is a Poisson random variable with mean 3. What is the probability that a 1-foot segment contains more than 4 defects?

Solution: This is

$$P(Y > 4) = \sum_{y=5}^{\infty} \frac{3^y e^{-3}}{y!} = 1 - \sum_{y=0}^4 \frac{3^y e^{-3}}{y!} \doteq .1847.$$

III. (10) (From an actuarial “P Exam” problem) The warranty on a machine specifies that it will be replaced at failure or at age 4, whichever occurs first. The machine’s age at failure, X , is a random variable with probability density function

$$f(x) = \begin{cases} 1/5 & \text{if } 0 < x < 5 \\ 0 & \text{otherwise} \end{cases}$$

Let Y be the age of the machine *at replacement*. What is $E(Y)$?

Solution: Note that X has a uniform distribution on the interval $[0, 5]$. The age of the machine at replacement is given by this function of X :

$$Y = \begin{cases} X & \text{if } 0 \leq X < 4 \\ 4 & \text{if } 4 \leq X \leq 5 \\ 0 & \text{otherwise} \end{cases}$$

So

$$E(Y) = \int_0^4 x \cdot \frac{1}{5} dx + \int_4^5 4 \cdot \frac{1}{5} dx = \frac{16}{10} + \frac{4}{5} = \frac{12}{5}.$$

IV.

A) (10) Let

$$g(x) = \begin{cases} \frac{1}{16}x^2 e^{x/2} & \text{if } x < 0 \\ 0 & \text{otherwise (that is, if } x \geq 0). \end{cases}$$

Show that g is a legal pdf for a random variable X .

Solution: $g(x) \geq 0$ for all x by inspection. Moreover

$$\int_{-\infty}^{\infty} g(x) dx = \int_{-\infty}^0 \frac{1}{16}x^2 e^{x/2} dx$$

Let $y = -x/2$. Then this integral becomes

$$\frac{1}{2} \int_0^{\infty} y^2 e^{-y} dy = \frac{1}{2} \Gamma(3) = 1.$$

Therefore all the requirements for a pdf are satisfied.

B) (10) Determine the moment-generating function of X . On which interval of t values is your formula valid?

Solution: The mgf is

$$\begin{aligned} E(e^{tX}) &= \int_{-\infty}^0 e^{tx} \cdot \frac{1}{16} x^2 e^{x/2} dx \\ &= \int_{-\infty}^0 \frac{1}{16} x^2 e^{(1+2t)x/2} dx \end{aligned}$$

Let $y = -(1 + 2t)x/2$ or $x = \frac{-2y}{1+2t}$. Making the substitution, we get (provided $1 + 2t > 0$),

$$\begin{aligned} &= \int_0^{\infty} \frac{1}{2(1+2t)^3} y^2 e^{-y} dy \\ &= \frac{1}{2(1+2t)^3} \Gamma(3) \\ &= \frac{1}{(1+2t)^3} \end{aligned}$$

In order for this to be valid, the improper integral must converge, and for that $1 + 2t > 0$, or $t > -1/2$.

V. Let Y_1, Y_2 be jointly continuous random variables with joint density

$$f(y_1, y_2) = \begin{cases} cy_1 & \text{if } 0 \leq y_1 \leq 1 \text{ and } y_1^2 \leq y_2 \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

A) (15) What is $P(Y_1 \leq Y_2)$?

Solution: The region where the density is nonzero is the region with $0 \leq y_1 \leq 1$, above the parabola $y_2 = y_1^2$, and below the line $y_2 = 1$. The value of c is determined as follows. We must have

$$\begin{aligned} 1 &= \int_0^1 \int_{y_1^2}^1 cy_1 dy_2 dy_1 \\ &= c \int_0^1 y_1 - y_1^3 dy_1 \\ &= c \left(\frac{1}{2} - \frac{1}{4} \right) \end{aligned}$$

So $c = 4$.

The condition $Y_1 \leq Y_2$ then defines the triangle with vertices at $(0, 0)$, $(1, 1)$, $(0, 1)$.

$$\begin{aligned} P(Y_1 \leq Y_2) &= \int_0^1 \int_{y_1}^1 4y_1 dy_2 dy_1 \\ &= \int_0^1 4y_1 - 4y_1^2 dy_1 \\ &= 2y_1^2 - \frac{4}{3}y_1^3 \Big|_0^1 \\ &= \frac{2}{3}. \end{aligned}$$

B) (15) Find the marginal density $f_1(y_1)$ and use it to compute $V(Y_1)$.

Solution: The marginal density is

$$f_1(y_1) = \begin{cases} \int_{y_1^2}^1 4y_1 \, dy_2 = 4y_1(1 - y_1^2) & \text{if } 0 \leq y_1 \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

Then

$$E(Y_1) = \int_0^1 4y_1^2 - 4y_1^4 \, dy_1 = \frac{4}{3} - \frac{4}{5} = \frac{8}{15}.$$

Similarly

$$E(Y_1^2) = \int_0^1 4y_1^3 - 4y_1^5 \, dy_1 = 1 - \frac{2}{3} = \frac{1}{3}.$$

So

$$V(Y_1) = \frac{1}{3} - \left(\frac{8}{15}\right)^2 = \frac{11}{225}.$$

Extra Credit. (10) What is $E(X)$ for the random variable X from question IV?

Solution: The answer is $E(X) = -6$. This can be computed either directly by evaluating $E(X) = \int_{-\infty}^0 x \cdot g(x) \, dx$, or more cleverly by using the mgf from the answer to IV B):

$$E(X) = \frac{d}{dt} \frac{1}{(1+2t)^3} \Big|_{t=0} = \frac{(-3)(2)}{(1+2 \cdot 0)^4} = -6.$$

(Note: the density $g(x)$ has the form of a gamma density with $\alpha = 3$ and $\beta = 2$, but reflected about $x = 0$. So the expected value should be the negative of the expected value of the corresponding standard gamma density: $\alpha\beta = 6$.)