

Mathematics 376 – Probability and Statistics II  
Solutions – Midterm Exam 2  
April 19, 2004

I.

A) We have

$$\begin{aligned} E(Y) &= \int_0^\theta y \cdot \frac{3y^2}{\theta^3} dy \\ &= \frac{3}{\theta^3} \frac{y^4}{4} \Big|_0^\theta \\ &= \frac{3\theta}{4} \end{aligned}$$

For the method of moments estimator, we set this equal to the sample mean (the first sample moment), and solve for  $\theta$ :

$$\frac{3\theta}{4} = \bar{Y} \Rightarrow \hat{\theta}_{MM} = \frac{4\bar{Y}}{3}$$

B) The likelihood function is

$$L(y_1, \dots, y_n | \theta) = \frac{3(y_1 \cdots y_n)^2}{\theta^{3n}}$$

if  $y_i \leq \theta$  for all  $i$ , and zero otherwise. This has no critical points (it's always decreasing as a function of  $\theta$ , for  $\theta > 0$ ) – the same is true for  $\ln(L)$ . Hence to maximize the likelihood, we want to take  $\theta$  as *small as possible*, consistent with the requirement that all the  $y_i$  must be less than or equal to  $\theta$ . This means that

$$\hat{\theta}_{ML} = Y_{(n)}$$

(the sample maximum).

II.

A) We have:

$$H_0 : p = .4$$

$$H_1 : p > .4$$

$$\text{Test statistic : } Z = \frac{Y/n - .4}{\sqrt{\frac{(.4)(.6)}{n}}}$$

$$RR : \{z : z > z_{.01}\} = \{z : z > 2.33\}$$

We use the value  $p = .4$  in the formula for the standard error of the estimator  $\hat{p} = Y/n$  because we are working under the assumption that  $H_0$  is true. See, for instance, the discussion in Example 10.6 on page 470 of the text.

B) Using the given data, our test statistic value is

$$z = \frac{1260/3000 - .4}{\sqrt{\frac{(.4)(.6)}{3000}}} \doteq 2.236$$

This is not in the rejection region, so we cannot reject the null hypothesis. In other words, this data is *not sufficient* to conclude that  $p$  is larger than .4, at the  $\alpha = .01$  level of significance.

C) The  $\hat{p}$ -value corresponding to the boundary point of the rejection region in A using  $n = 3000$  is  $.4 + (2.33)\sqrt{\frac{(.4)(.6)}{3000}} = .42084$ . We want

$$\begin{aligned}\beta &= P(z \notin RR | p = .43) \\ &= P(Y/3000 < .42084 | p = .43) \\ &= P\left(z < \frac{.42084 - .43}{\sqrt{\frac{(.43)(.57)}{3000}}}\right) \\ &= P(z < -1.01) \\ &= P(z > 1.01) \doteq .1562\end{aligned}$$

(Note that we use the value  $p = .43$  here in the formula for the standard error of the estimator  $Y/n$  because we are now working under the assumption that  $p$  has actually increased from .4 to calculate the Type II error probability.)

### III.

A) Since the sample size is small ( $7 < 30$ ), we use a one-tail  $t$ -test for a single mean here. The null hypothesis is that  $\mu = \mu_0 = 172$ , the alternative is that  $\mu < 172$ . We compute  $\bar{x} = 167.86$  and  $S^2 = \frac{1}{6} \sum_{i=1}^6 (x_i - \bar{x})^2 = 43.81$ . Then our test statistic is

$$t = \frac{\bar{x} - \mu_0}{S/\sqrt{n}} = \frac{167.86 - 172}{\sqrt{43.81/7}} \doteq -1.655$$

The rejection region for the lower-tail  $t$ -test is

$$RR = \{t : t < -t_{.1}(6)\} = \{t : t < -1.44\}$$

from the  $t$ -table. Since  $t$  is in the rejection region, we reject the null hypothesis. In other words, there is sufficient evidence to say that  $\mu < 172$  (at the  $\alpha = .1$  level of significance).

B) To test  $\sigma^2 = 40$  against  $\sigma^2 \neq 40$ , we want a two-tailed  $\chi^2$ -test. The test statistic is

$$\chi^2 = 6S^2/\sigma^2 = 6(43.81)/40 = 6.57135$$

The rejection region is

$$RR = \{\chi^2 < \chi_{.95}^2(6)\} \cup \{\chi^2 > \chi_{.05}^2(6)\} = \{\chi^2 < 1.63539\} \cup \{\chi^2 > 12.5916\}$$

Our value 6.57135 is not in the union of these two intervals, so we do not have sufficient evidence to reject the null hypothesis here.

IV.

- A) Using the example of a small sample test on a single mean (like III A above), the test statistic

$$\frac{\bar{X} - \mu_0}{S/\sqrt{n}}$$

has a  $t$ -distribution with  $n - 1$  degrees of freedom. For small  $n$  ( $n < 30$ ), this is different enough from a standard normal (more probability in the tails) that the  $z$ -test would lead to rejection regions that are “too large”. For example, with an upper tail test at the  $\alpha = .1$  level and  $n = 7$ ,  $t_{.1}(6) = 1.44$ , but  $z_{.1} = 1.28$ .

- B) The  $p$ -value for the two-tailed test would be twice as large as the  $p$ -value for the one-tail test. For example, in III A above, the  $p$ -value for the one-tail test is  $< .1$  and  $> .05$  from the book’s table. Exact value using Maple is approximately .0745. If we were doing a two-tail  $t$ -test, the value 1.655 would be  $t_{\alpha/2}(6)$  if  $\alpha/2 = .0745$ , so  $\alpha = 2(.0745) = .1490$  is the  $p$ -value of the test.