

Probability + Statistics

Final Review Answers - Part II

18. w/ 1000 samples, binomial dist. is basically normal

$$Z = \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

$$= \frac{5/1000 - 0.01}{\sqrt{\frac{0.01(1-0.01)}{1000}}}$$

$$= -2.24$$

$$P(Z \leq -2.24)$$

$$= \Phi(-2.24)$$

$$= 1 - \Phi(2.24)$$

$$= 1 - 0.9875$$

$$= 0.0125$$

Redone
@ end of
solutions

19. w/ 100 patients, dist is normal

$$\hat{p} = \frac{y}{n}$$

$$\hat{p} = 75/100$$

$$\hat{p} = 0.75$$

$$Z = \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

$$= \frac{0.75 - 0.667}{\sqrt{\frac{0.667(1-0.667)}{100}}}$$

$$= 1.9168$$

$$P(Z \geq 1.9168)$$

$$= 1 - \Phi(1.9168)$$

$$= 1 - 0.9726$$

$$= 0.0274$$

P = likelihood a
Patient is female
= 0.667

$$Z_{0.05} = 1.645$$

$$Z_{0.05} = \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

$$1.645 = \frac{0.7 - 0.667}{\sqrt{\frac{0.667(1-0.667)}{n}}}$$

$$n = 522$$

~~$$Z_{0.05} = \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

$$1.645 = \frac{0.7 - 0.667}{\sqrt{\frac{0.667(1-0.667)}{n}}}$$~~

20. 8 days = 192 hours

$$\theta = 96$$

$$P(3 \text{ failures before } t=192)$$

$$= P(1 \text{ failure before } t=192)^3$$

$$= \left[\int_0^{192} \frac{1}{96} e^{-t/96} dt \right]^3$$

$$= 0.646$$

Redone
@ End
of Sol'n's

~~$$P(\text{failure}) = \int_0^{192} \frac{1}{96} e^{-t/96} dt$$

$$= 0.8647$$~~

21. $t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$

$$= \frac{50 - 52}{3/\sqrt{10}}$$

$$= -2.1082$$

$$P(t(9) \leq -2.1082)$$

$$= 1 - P(t(9) \leq 2.1082)$$

~~between~~ between 0.05 and 0.025

22. ~~$$f(x) = \frac{\theta}{2} e^{-\theta|x|}$$

$$\ln f(x) = \ln(\theta/2) - \theta|x|$$~~

$$L(\theta) = \prod_{i=1}^n \frac{\theta}{2} e^{-\theta|x_i|}$$

$$= \left(\frac{\theta}{2}\right)^n e^{-\theta \sum_{i=1}^n |x_i|}$$

$$\ln(L(\theta)) = n \ln(\theta/2) - \theta \sum_{i=1}^n |x_i|$$

$$= n \ln \theta - n \ln 2 - \theta \sum_{i=1}^n |x_i|$$

$$\frac{d}{d\theta} [\ln(L(\theta))] = \frac{n}{\theta} - \sum_{i=1}^n |x_i|$$

$$0 = \frac{n}{\theta} - \sum_{i=1}^n |x_i|$$

$$\sum_{i=1}^n |x_i| = \frac{n}{\theta}$$

$$\theta = n / \sum_{i=1}^n |x_i|$$

$$\frac{d^2}{d\theta^2} \ln(L) = -\frac{n}{\theta^2}$$

always positive

$$= -\frac{n}{\theta^2} \left(\sum_{i=1}^n |x_i| \right)^2$$

always positive

n^2

∴ 2nd derivative is always negative,

so $\theta = \frac{n}{\sum_{i=1}^n |x_i|}$ is

a maximum

23. a) assuming duration is normally distributed (when averaged over 30 patients) this is Kosher b/c of the central limit theorem

$$Z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$$

$$= \frac{7 - 8}{4/\sqrt{30}}$$

$$= -1.3693$$

$$P(Z \leq -1.3693)$$

$$= 1 - \Phi(1.3693)$$

$$= 1 - 0.9147$$

$$= 0.0853$$

$$b) Z_2 = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$$

$$= \frac{9-8}{4/\sqrt{30}}$$

$$= 1.3693$$

$$P(-1.3693 \leq Z \leq 1.3693)$$

$$= \Phi(1.3693) - (1 - \Phi(1.3693))$$

$$= 2\Phi(1.3693) - 1$$

$$= 2 \cdot 0.9147 - 1$$

$$= 0.8294$$

c) Assuming \bar{X} is normally distributed b/c underlying distribution is normal

24. a) Assuming normal distribution and that the two distributions are independent

$$b) t = \frac{\bar{x} - \bar{y}}{\sqrt{\frac{(n-1)s_x^2 + (m-1)s_y^2}{n+m-2} \left(\frac{1}{n} + \frac{1}{m}\right)}}$$

$$= \frac{58-53}{\sqrt{\frac{(54-1)15^2 + (30-1)(13)^2}{54+30-2} \left(\frac{1}{54} + \frac{1}{30}\right)}}$$

$$= 1.53$$

$$\text{Using } r = n+m-2$$

$$= 82$$

$$P(Z > 1.53) = .06$$

Since this is a two-tailed test

$$p\text{-value} = P(Z > 1.53) * 2 = 0.12$$

c. Fail to reject

distr is basically normal

$$25. a) t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$$

$$= \frac{47.1 - 48}{\sqrt{4.7/12}}$$

$$= -1.438$$

Critical region $|t| \geq t_{\alpha/2}(r)$

$$|t| \geq t_{0.05}(11)$$

$$|t| \geq 1.796$$

Fail to reject

b) It's between ~~0.10~~ $0.10 \leq \frac{\alpha}{2} \leq 0.05$

So p-value btw 0.2 and 0.1

c) ~~111~~ $\bar{X} \pm t_{\alpha/2}(r) \cdot s_x / \sqrt{n}$
 $47.1 \pm 1.796 \cdot \sqrt{4.7/12}$
 $[45.976, 48.224]$

~~If we repeated the experiment many times, the true mean would be in our interval 90% of the time.~~

If ~~not~~ we repeated the experiment many times, 90% of the CI's generated would contain μ (the true population mean)

26. critical region:

$$\frac{s_1^2}{s_2^2} \geq F_{\alpha}(n_1 - 1, n_2 - 1) \quad (\text{Using } \alpha = 0.05)$$

$$\frac{1.04}{0.51} \geq 1.98 \quad \text{reject } H_0$$

$$2.029 \geq 1.98$$

27. $P(\text{Flu} | \text{Rash}) = \frac{P(\text{Rash} | \text{Flu}) \cdot P(\text{Flu})}{P(\text{Rash})}$
 $= \frac{(0.05)(0.8)}{(0.2)(0.9) + (0.8)(0.05)}$
 $= 0.18$

$$P(\text{Rash}) = P(\text{M and Rash}) + P(\text{F \& R})$$
$$= 0.2 \cdot 0.9 + 0.8 \cdot 0.05$$

28. $P(\text{Tails}) = 0.5$

game ends on first tails

$$p = 0.5$$

Average duration = $\frac{1}{p} = 2 \text{ flips}$

It's a geometric distribution

$$\text{Expected winnings} = \sum_{i=1}^{\infty} x_i f(x_i)$$

$$= \sum_{i=1}^{\infty} 2^{i-1} \left(\frac{1}{2}\right)^i$$

$$= \sum_{i=1}^{\infty} 2^{-1}$$

~~It~~ does not converge

Expected value
is infinite!

29. $F(x) = \int_{-\infty}^x f(t) dt$

$$= \int_{-\infty}^x \frac{1}{2} \alpha e^{-\alpha |t|} dt$$

$$= \frac{1}{2} \alpha \int_{-\infty}^x e^{-\alpha |t|} dt$$

$$= \frac{1}{2} \alpha \left[\int_{-\infty}^0 e^{\alpha x} dx + \int_0^x e^{-\alpha x} dx \right]$$

4) $F(x) = \int_{-\infty}^x f(t) dt$

$$= \int_{-\infty}^x \frac{1}{2} \alpha e^{-\alpha |t|} dt$$

$$= \frac{1}{2} \alpha \int_{-\infty}^x e^{-\alpha |t|} dt$$

$x \leq 0$

$$= \frac{1}{2} \alpha \int_{-\infty}^x e^{\alpha t} dt$$

$$= \frac{1}{2} \alpha \left[\frac{1}{\alpha} e^{\alpha t} \right]_{-\infty}^x$$

$$= \frac{1}{2} \alpha \left[\frac{1}{\alpha} e^{\alpha x} \right]$$

$$= \frac{1}{2} e^{\alpha x}$$

$x > 0$

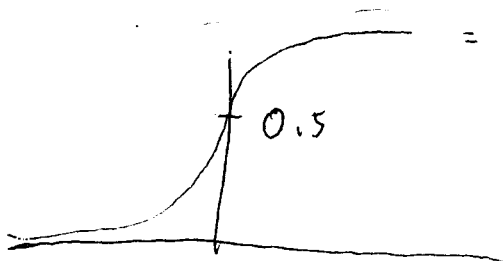
$$\frac{1}{2} \alpha \left[\int_{-\infty}^0 e^{\alpha t} dt + \int_0^x e^{-\alpha t} dt \right]$$

$$\frac{1}{2} \alpha \left[\frac{1}{\alpha} + \left[-\frac{1}{\alpha} e^{-\alpha t} \right]_0^x \right]$$

$$\frac{1}{2} \alpha \left[\frac{1}{\alpha} - \frac{1}{\alpha} e^{-\alpha x} + \frac{1}{\alpha} \right]$$

$$= \frac{1}{2} \alpha \left[\frac{2}{\alpha} - \frac{1}{\alpha} e^{-\alpha x} \right]$$

$$= 1 - \frac{1}{2} e^{-\alpha x}$$



b) ~~$E[X] = \int_{-\infty}^{\infty} x f(x) dx$~~
 ~~$= \int_{-\infty}^{\infty} x \cdot \frac{1}{2} \alpha e^{-\alpha|x|} dx$~~

Since $F(0) = 0.5$, 0 is the mean of X

$$\text{Var}(X) = E[X^2] - \mu^2$$

$$= \int_{-\infty}^{\infty} x^2 \cdot \frac{1}{2} \alpha e^{-\alpha|x|} dx - 0$$

$$= \frac{1}{2} \int_{-\infty}^0 x^2 \alpha e^{-\alpha|x|} dx + \frac{1}{2} \int_0^{\infty} x^2 \alpha e^{-\alpha|x|} dx$$

$$= \frac{1}{2} \int_0^{\infty} x^2 \alpha e^{-\alpha x} dx + \frac{1}{2} \int_0^{\infty} x^2 \alpha e^{-\alpha x} dx$$

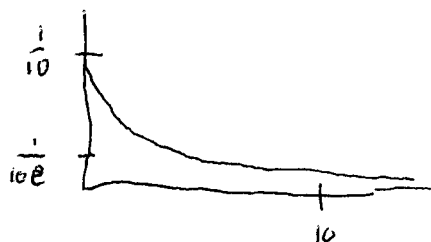
$$= \frac{1}{2} \cdot \left(\frac{2}{\alpha^2} \right) + \frac{1}{2} \cdot \left(\frac{2}{\alpha^2} \right)$$

$$= 2/\alpha^2$$

integration by parts
 (see page 120
 in textbook)

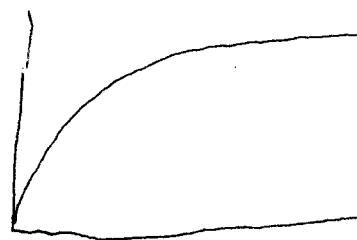
$$\therefore \text{Std}(X) = \sqrt{2}/\alpha$$

30. a) $f(x) = \frac{1}{\theta} e^{-x/\theta}$
 $= \frac{1}{10} e^{-x/10}$



$$F(x) = 1 - e^{-x/\theta}$$

$$= 1 - e^{-x/10}$$



$$P(X \geq 10)$$

$$= 1 - P(X < 10)$$

$$= 1 - F(10)$$

$$= 1 - [1 - e^{-10/10}]$$

$$= 1 - [1 - \frac{1}{e}]$$

$$= \frac{1}{e}$$

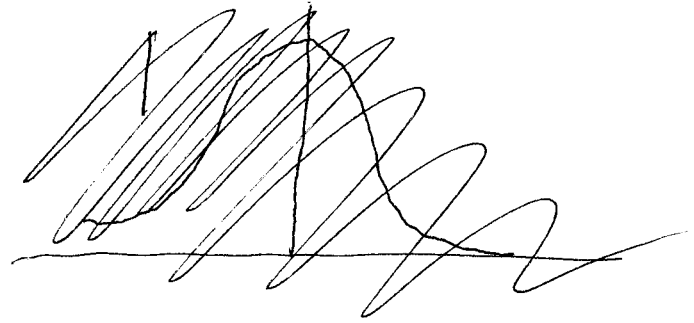
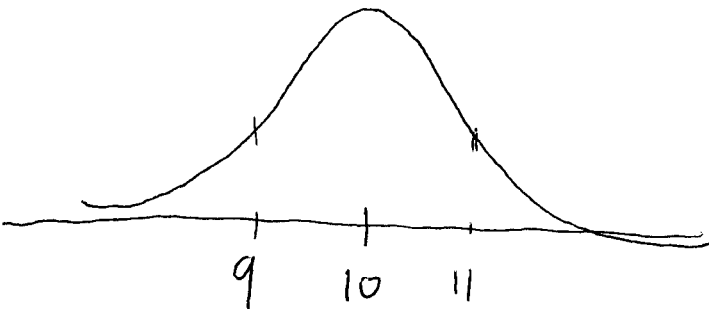
b) w/ 100 items, \bar{X} will be basically normal
b/c central limit thm

$$\mu = 10$$

$$\sigma_{\bar{x}} = \sigma_x / \sqrt{n} = 10 / \sqrt{100} = 1$$

~~Different~~

$$\bar{X} \sim N(\mu=10, \sigma=1)$$



31. (See problem 23)

32. (See problem 25)

33. a) Assuming both distributions are normal. This is kosher
b/c of the central limit theorem

Assuming both distributions are independent b/c they are composed
of different individuals

$$b) \hat{p}_1 = 337/500 = 0.674$$

$$\hat{p}_2 = 215/500 = \del{0.430} 0.430$$

$$n_1 = 500$$

$$n_2 = 500$$

$$Z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}}$$

$$= 8.00$$

$$P \text{ value} \approx 0$$

$$c) \hat{p}_1 - \hat{p}_2 \pm Z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

$$Z_{\alpha/2} = 1.96$$

$$[0.18, 0.30]$$

d) Reject

34. a) Test statistic: $\chi^2 = \frac{(n-1)s^2}{\sigma_0^2}$ $df: n-1 = 10-1 = 9$

b) $H_0: \text{range} = 4\sigma$ $H_a: \sigma > 10$
 $40 = 4\sigma_0$
 $\sigma_0 = 10$

c) $\chi^2 = \frac{(n-1)s^2}{\sigma_0^2}$
 $= \frac{(9)(195)}{10^2}$
 $= 17.55$

d) $\chi^2 \geq \chi_{\alpha}^2(n-1)$
 $\geq \chi_{0.05}^2(9)$
 ≥ 16.92

e) ~~$[0, \frac{(n-1)s^2}{\chi_{\alpha}^2(n-1)}]$~~ ~~$[0, \frac{(n-1)s^2}{\chi_{\alpha}^2(n-1)}]$~~ $[\frac{(n-1)s^2}{\chi_{\alpha}^2(n-1)}, \infty)$
 ~~$[0, \frac{9 \cdot 195}{16.92}]$~~ ~~$[0, \frac{9 \cdot 195}{16.92}]$~~ $[\frac{9 \cdot 195}{16.92}, \infty)$
 ~~$[0, 103.72]$~~ ~~$[0, 103.72]$~~ $[103.72, \infty)$
 ~~$[\frac{(n-1)s^2}{\chi_{\alpha}^2(n-1)}, \infty)$~~ ~~$[\frac{(n-1)s^2}{\chi_{\alpha}^2(n-1)}, \infty)$~~ $[\frac{(n-1)s^2}{\chi_{\alpha}^2(n-1)}, \infty)$
 ~~$[\frac{9 \cdot 195}{16.92}, \infty)$~~ ~~$[\frac{9 \cdot 195}{16.92}, \infty)$~~ $[\frac{9 \cdot 195}{16.92}, \infty)$
 ~~$[103.72, \infty)$~~ ~~$[103.72, \infty)$~~ $[103.72, \infty)$

f) REJECT

$$18. P(\text{Errors} \leq 5)$$

$$= P(0 \text{ Error}) + P(1 \text{ Error}) + P(2 \text{ Error}) + P(3 \text{ Error}) + P(4 \text{ Error}) + P(5 \text{ Error})$$

$$= \binom{1000}{0} (0.99)^{1000} (0.01)^0 + \binom{1000}{1} (0.99)^{999} (0.01)^1 + \binom{1000}{2} (0.99)^{998} (0.01)^2$$

$$+ \binom{1000}{3} (0.99)^{997} (0.01)^3 + \binom{1000}{4} (0.99)^{996} (0.01)^4 + \binom{1000}{5} (0.99)^{995} (0.01)^5$$

$$20. P(\text{Single computer fails before 192 hours})$$

$$= \int_0^{192} \frac{1}{96} e^{-t/96} dt$$

$$= 0.8647$$

$$P(\text{at least 3 out of 5 computers fail})$$

~~$$= P(0 \text{ fail}) + P(1 \text{ fail}) + P(2 \text{ fail}) + P(3 \text{ fail})$$~~
~~$$= \binom{5}{0} (1-0.8647)^5 (0.8647)^0 + \binom{5}{1} (1-0.8647)^4 (0.8647)^1$$~~
~~$$+ \binom{5}{2} (1-0.8647)^3 (0.8647)^2 + \binom{5}{3} (1-0.8647)^2 (0.8647)^3$$~~

$$= P(5 \text{ fail}) + P(4 \text{ fail}) + P(3 \text{ fail})$$

$$= \binom{5}{5} (0.8647)^5 (1-0.8647)^0 + \binom{5}{4} (0.8647)^4 (1-0.8647)^1$$

$$+ \binom{5}{3} (0.8647)^3 (1-0.8647)^2$$

